# Package 'reverseR'

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<b>Description</b> Tests linear regressions for significance reversal through leave-one(multiple)-out and shifting/addition of response values. The paradigm of the package is loosely based on the somewhat forgotten ``dfstat" criterion (Belsley, Kuh & Welsch 1980 <doi:10.1002 0471725153.ch2="">), which tests influential values in linear models from their effect on statistical inference, i.e. changes in p-value.</doi:10.1002>
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Influence plots

Several diagnostic plots for checking p-value influencers

# **Description**

Seven different plot types that visualize *p*-value influencers.

- 1. lmPlot: plots the linear regression, marks the influencer(s) in red and displays trend lines for the full and leave-one-out (LOO) data set (black and red, respectively).
- 2. pvalPlot: plots the p-values for each LOO data point and displays the values as a full model/LOO model plot, together with the alpha border as defined in lmInfl.
- 3. inflPlot: plots dfbeta for slope, dffits, covratio, cooks.distance, leverage (hatvalues) and studentized residuals (rstudent) against the  $\Delta p$ -value. Herewith, changes in these six parameters can be compared to the effect on the corresponding drop/rise in p-value. The plots include vertical boundaries for threshold values as defined in the literature under 'References'.
- 4. slsePlot: plots all LOO-slopes and their standard errors together with the corresponding original model values and a t-value border as calculated by  $Q_t(1-\frac{\alpha}{2},n-2)$ . LOO of points on the right of this border result in a significant model, and *vice versa*.
- 5. threshPlot: plots the output of lmThresh, i.e. the regression plot including confidence/prediction intervals, as well as for each response value  $y_i$  the region in which the model is significant (green). This is tested for either i)  $y_i$  that are shifted into this region (newobs = FALSE in lmThresh) or ii) when a new observation  $y_i$  is added (newobs = TRUE in lmThresh). In the latter case, it is informative if this region resides within the prediction interval (dashed line), indicating that a future additional measurement at  $x_i$  might reverse the significance statement.
- 6. multPlot: plots the output of lmMult as a point cloud of p-values for each 1...max sample removals and n combinations. All combinations for which the sample removal resulted in a significance reversal are colored in red, the percentages of these are given on top of the plot.
- 7. stabPlot: for single (to be selected) response values from the output of lmThresh, this function displays the region of significance reversal within the surrounding prediction interval. The probability of a either shifting the response value (if lmThresh(..., newobs = FALSE)) or of including a future (measurement) point (if lmThresh(..., newobs = TRUE)) to reverse the significance is shown as the integral between the "end of significance region" (eosr) and the nearest prediction interval boundary.

**NOTE**: The visual display should always be supplemented with the corresponding stability analysis.

## Usage

```
lmPlot(infl, ...)
pvalPlot(infl, ...)
inflPlot(infl, ...)
slsePlot(infl, ...)
threshPlot(thresh, bands = FALSE, ...)
multPlot(mult, log = FALSE, ...)
stabPlot(stab, which = NULL, ...)
```

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# **Arguments**

infl an object obtained from lmInfl.

thresh an object obtained from lmThresh.

stab an object obtained from using stability on an lmThresh output.

bands logical. If TRUE, plots the confidence and prediction bands.

mult an object obtained from lmMult.

log should the p-values be displayed on a logarithmic y-axis?

which which response value should be shown in stabPlot?

... other plotting parameters.

#### Value

The corresponding plot.

### Note

Cut-off values for the different influence measures are those defined in Belsley, Kuh E & Welsch (1980):

```
dfbeta slope: |\Delta\beta 1_i|>2/\sqrt{n} dffits: |\mathrm{dffits}_i|>2\sqrt{2/n} covratio: |\mathrm{covr}_i-1|>3k/n Cook's D: D_i>Q_F(0.5,k,n-k) leverage: h_{ii}>2k/n studentized residual: t_i>Q_t(0.975,n-k-1)
```

#### Author(s)

Andrej-Nikolai Spiess

SAGE Publishing, 3rd ed, 2016.

# References

```
Regression diagnostics: Identifying influential data and sources of collinearity. Belsley DA, Kuh E, Welsch RE. John Wiley, New York (1980).

Applied Regression Analysis: A Research Tool.
Rawlings JO, Pantula SG, Dickey DA.
Springer; 2nd Corrected ed. 1998. Corr. 2nd printing 2001.

Applied Regression Analysis and Generalized Linear Models.
Fox J.
```

```
## See Examples in 'lmInfl', 'lmThresh' and 'lmMult'.
```

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1mExact Create random values that deliver linear regressions rameters	s with exact pa-
---	------------------

# **Description**

Takes self-supplied x/y values or x/random values and transforms these as to deliver linear regressions  $y = \beta_0 + \beta_1 x + \varepsilon$  (with potential replicates) with either

```
1) exact slope \beta_1 and intercept \beta_0,
```

- **2)** exact *p*-value and intercept  $\beta_0$ , or
- 3) exact  $R^2$  and intercept  $\beta_0$ .

Intended for testing and education, not for cheating ! ;-)

# Usage

## **Arguments**

X	the predictor values.
У	NULL. A possible vector of $y$ values with length(x).
ny	the number of replicate response values per predictor value.
intercept	the desired intercept $\beta_0$ .
slope	the desired slope $\beta_1$ .
error	if a single value, the standard deviation $\sigma$ for sampling from a normal distribution, or a user-supplied vector of length x with random deviates.
seed	the random generator seed for reproducibility.
pval	the desired <i>p</i> -value of the slope.
rsq	the desired $R^2$ .
plot	logical. If TRUE, the linear regression is plotted.
verbose	logical. If TRUE, a summary is printed to the console.
• • •	other arguments to lm or plot.

# **Details**

For case 1), the error values are added to the exact  $(x_i, \beta_0 + \beta_1 x_i)$  values, the linear model  $y_i = \beta_0 + \beta_1 x_i + \varepsilon$  is fit, and the residuals  $y_i - \hat{y_i}$  are re-added to  $(x_i, \beta_0 + \beta_1 x_i)$ .

For case 2), the same as in 1) is conducted, however the slope delivering the desired p-value is found by an optimizing algorithm.

Finally, for case 3), a QR reconstruction, rescaling and refitting is conducted, using the code found

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under 'References'.

If y is supplied, changes in slope, intercept and p-value will deliver the sames residuals as the linear regression through x and y. A different  $R^2$  will change the response value structure, however.

#### Value

A list with the following items:

1m the linear model of class 1m.

x the predictor values.

y the (random) response values.

summary the model summary for quick checking of obtained parameters.

Using both x and y will give a linear regression with the desired parameter values when refitted.

## Author(s)

Andrej-Nikolai Spiess

#### References

For method 3):

http://stats.stackexchange.com/questions/15011/generate-a-random-variable-with-a-defined-correlation-to-an-existing-variable.

```
## No replicates, intercept = 3, slope = 0.2, sigma = 2, n = 20.
res1 <- lmExact(x = 1:20, ny = 1, intercept = 3, slope = 2, error = 2)
## Same as above, but with 3 replicates, sigma = 1, n = 20.
res2 <- lmExact(x = 1:20, ny = 3, intercept = 3, slope = 2, error = 1)
## No replicates, intercept = 2 and p-value = 0.025, sigma = 3, n = 50.
## => slope = 0.063
res3 <- lmExact(x = 1:50, ny = 1, intercept = 2, pval = 0.025, error = 3)
## 5 replicates, intercept = 1, R-square = 0.85, sigma = 2, n = 10.
## => slope = 0.117
res4 <- lmExact(x = 1:10, ny = 5, intercept = 1, rsq = 0.85, error = 2)
## Heteroscedastic (magnitude-dependent) noise.
error \leftarrow sapply(1:20, function(x) rnorm(3, 0, x/10))
res5 <- lmExact(x = 1:20, ny = 3, intercept = 1, slope = 0.2,
                error = error)
## Supply own x/y values, residuals are similar to an
## initial linear regression.
X \leftarrow c(1.05, 3, 5.2, 7.5, 10.2, 11.7)
```

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```
set.seed(123) Y \leftarrow 0.5 + 2 * X + rnorm(6, 0, 2) res6 \leftarrow lmExact(x = X, y = Y, intercept = 1, slope = 0.2) all.equal(residuals(lm(Y \sim X)), residuals(res6$lm))
```

lmInf1

Checks and analyzes leave-one-out (LOO) p-values in linear regression

## **Description**

This function calculates leave-one-out (LOO) p-values for all data points and identifies those resulting in "significance reversal", i.e. in the p-value of the model's slope traversing the user-defined  $\alpha$ -level.

## Usage

```
lmInfl(model, alpha = 0.05, method = c("pearson", "spearman"), verbose = TRUE, ...)
```

## **Arguments**

model the linear model of class lm.

alpha the  $\alpha$ -level to use as the threshold border.

method select either parametric ("pearson") or rank-based ("spearman") statistics.

verbose logical. If TRUE, results are displayed on the console.

... other arguments to lm.

### **Details**

The algorithm

- 1) calculates the *p*-value of the full model (all points),
- 2) calculates a LOO-p-value for each point removed,
- 3) checks for significance reversal in all data points and
- 4) returns all models as well as classical influence.measures with LOO-p-values,  $\Delta p$ -values, slopes and standard errors attached.

If method = "spearman", p-values are based on Spearman Rank correlation, and the values given in the last column of the result matrix are Spearman's  $\rho$ .

The idea of *p*-value influencers was first introduced by Belsley, Kuh & Welsch, and described as an influence measure pertaining directly to the change in *t*-statistics, that will "show whether the conclusions of hypothesis testing would be affected", termed **dfstat** in [1, 2, 3] or **dfstud** in [4]:

$$dfstat_{ij} \equiv \frac{\hat{\beta}_{j}}{s\sqrt{(X'X)_{jj}^{-1}}} - \frac{\hat{\beta}_{j(i)}}{s_{(i)}\sqrt{(X'_{(i)}X_{(i)})_{jj}^{-1}}}$$

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where  $\hat{\beta}_j$  is the *j*-th estimate, *s* is the residual standard error, *X* is the design matrix and (*i*) denotes the *i*-th observation deleted.

**dfstat**, which for the regression's slope  $\beta_1$  is the difference of *t*-statistics

$$\Delta t = t_{\beta 1} - t_{\beta 1(i)} = \frac{\beta_1}{\text{s.e.}(\beta_1)} - \frac{\beta_1(i)}{\text{s.e.}(\beta_1(i))}$$

is inextricably linked to the changes in p-value  $\Delta p$ , calculated from

$$\Delta p = p_{\beta 1} - p_{\beta 1(i)} = 2 \left( 1 - P_t(t_{\beta 1}, \nu) \right) - 2 \left( 1 - P_t(t_{\beta 1(i)}, \nu - 1) \right)$$

where  $P_t$  is the Student's t cumulative distribution function with  $\nu$  degrees of freedom, and where significance reversal is attained when  $\alpha \in [p_{\beta 1}, p_{\beta 1(i)}]$ . Interestingly, in linear regression the seemingly mandatory check of the influence of single data points on statistical inference is living in oblivion: apart from [1-4], there is, to the best of our knowledge, no reference to **dfstat** or  $\Delta p$  in current literature on influence measures.

The influence output also includes the more recent Hadi's measure (column "hadi"):

$$H_i = \frac{p_{ii}}{1 - p_{ii}} + \frac{k}{1 - p_{ii}} \frac{d_i^2}{(1 - d_i^2)}$$

where  $p_{ii}$  are the diagonals of the hat matrix (leverages), k=2 in univariate linear regression and  $d_i = e_i/\sqrt{\text{SSE}}$ .

#### Value

A list with the following items:

origModel	the original model with all data points.
finalModels	a list of final models with the influencer(s) removed.
infl	a matrix with the original data, classical influence.measures, studentized residuals, leverages, LOO- $p$ -values, LOO-slopes/intercepts and their $\Delta$ 's, LOO-standard errors and $R^2$ s.
sel	a vector with the influencers' indices.
alpha	the selected $\alpha$ -level.
origP	the original model's <i>p</i> -value.
stab	the stability measure, see stability.

#### Author(s)

Andrej-Nikolai Spiess

## References

#### For dfstat / dfstud:

1. Regression diagnostics: Identifying influential data and sources of collinearity. Belsley DA, Kuh E, Welsch RE.

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```
John Wiley, New York, USA (2004).
```

2. Econometrics, 5ed.

Baltagi B.

Springer-Verlag Berlin, Germany (2011).

3. Growth regressions and what the textbooks don't tell you.

Temple J.

Bull Econom Res, 52, 2000, 181-205.

4. Robust Regression and Outlier Detection.

Rousseeuw PJ & Leroy AM.

John Wiley & Sons, New York, NY (1987).

#### Hadi's measure:

A new measure of overall potential influence in linear regression.

Hadi AS.

Comp Stat & Data Anal, 14, 1992, 1-27.

```
## Example #1 with single influencers and insignificant model (p = 0.115).
## Removal of #18 results in p = 0.0227!
set.seed(123)
a <- 1:20
b < -5 + 0.08 * a + rnorm(20, 0, 1)
LM1 <- lm(b \sim a)
res1 <- lmInfl(LM1)</pre>
lmPlot(res1)
pvalPlot(res1)
inflPlot(res1)
slsePlot(res1)
stability(res1)
## Example #2 with multiple influencers and significant model (p = 0.0269).
## Removal of #2, #17, #18 or #20 result in crossing p = 0.05!
set.seed(125)
a <- 1:20
b < -5 + 0.08 * a + rnorm(20, 0, 1)
LM2 <- lm(b \sim a)
res2 <- lmInfl(LM2)</pre>
lmPlot(res2)
pvalPlot(res2)
inflPlot(res2)
slsePlot(res2)
stability(res2)
## Large Example #3 with top 10 influencers and significant model (p = 6.72E-8).
## Not possible to achieve a crossing of alpha with any point despite strong noise.
set.seed(123)
```

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```
a <- 1:100
b < -5 + 0.08 * a + rnorm(100, 0, 5)
LM3 <- lm(b \sim a)
res3 <- lmInfl(LM3)</pre>
lmPlot(res3)
stability(res3)
## Example #4 with replicates and single influencer (p = 0.114).
## Removal of #58 results in p = 0.039.
set.seed(123)
a <- rep(1:20, each = 3)
b < -5 + 0.08 * a + rnorm(20, 0, 2)
LM4 <- lm(b \sim a)
res4 <- lmInfl(LM4)</pre>
lmPlot(res4)
pvalPlot(res4)
inflPlot(res4)
slsePlot(res4)
stability(res4)
## As Example #1, but with weights.
## Removal of #18 results in p = 0.04747.
set.seed(123)
a <- 1:20
b < -5 + 0.08 * a + rnorm(20, 0, 1)
LM5 <- lm(b \sim a, weights = 1:20)
res5 <- lmInfl(LM5)</pre>
lmPlot(res5)
stability(res5)
```

lmMult

Checks and analyzes leave-multiple-out (LMO) p-values in linear regression

### **Description**

This function calculates leave-multiple-out (LMO) p-values for an increasing number of data points and identifies those resulting in "significance reversal" of the model, i.e. in the slope's p-value traversing the user-defined  $\alpha$ -level.

# Usage

# Arguments

model the linear model of class 1m.

max the maximum number of points to eliminate.

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n the number of samples to draw for each 1...max.

alpha the  $\alpha$ -level to use as the threshold border.

method select either parametric ("pearson") or rank-based ("spearman") statistics.

verbose logical. If TRUE, results for each 1...max will be printed to the console.

#### **Details**

The algorithm

- 1) calculates the *p*-value of the full model (all data points),
- 2) calculates a LMO-p-value for all n sampled groups of 1...max points removed,
- 3) checks for significance reversal in the resulting model and
- 4) returns all n samples and the corresponding *p*-values.

#### Value

A list with the following items:

sample a matrix with all max \* n iterations, where a 1 indicates the left-out sample(s),

as well as the corresponding p-values and group.

stat for each 1...max LMO's, the percentage of model significance reversers.

### Author(s)

Andrej-Nikolai Spiess

```
## Example with single influencers and insignificant model (p = 0.115).
set.seed(123)
a <- 1:20
b < -5 + 0.08 * a + rnorm(20, 0, 1)
LM1 <- lm(b \sim a)
res1 <- lmMult(LM1)</pre>
multPlot(res1)
stability(res1)
## Large example with 100 data points and highly significant model (p = 6.72E-8).
## No significance reversal up to the elimination of 20 points.
set.seed(123)
a <- 1:100
b < -5 + 0.08 * a + rnorm(100, 0, 5)
LM2 \leftarrow lm(b \sim a)
res2 <- lmMult(LM2, max = 20)
multPlot(res2)
stability(res2)
```

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lmThresh	Finds and analyzes significance reversal regions for each response value

### **Description**

This function finds (by iterating through a grid of values for each response) the approximate response value range(s) in which the regression is significant (when inside) or not (when outside), as defined by alpha. Here, two scenarios can be tested: i) if newobs = FALSE (default), the model's significance is tested by shifting  $y_i$  along the search grid. If newobs = TRUE,  $y_i$  is kept fixed and a new observation  $y_{2i}$  is added and shifted along the search grid. Hence, this function tests the regression for the sensitivity of being reversed in its significance through minor shifting of the original or added response values, as opposed to the effect of point removal (lmInfl).

## Usage

#### **Arguments**

model	the linear model of class 1m.
factor	a factor for the initial search grid. See 'Details'.
alpha	the $\alpha$ -level to use as the threshold border.
method	select either parametric ("pearson") or rank-based ("spearman") statistics.
steps	the number of steps within the search range. See 'Details'.
newobs	logical. Should the significance region for each $y_i$ be calculated from shifting $y_i$ or from keeping $y_i$ fixed and adding a new observation $y2_i$ ?
	other arguments to future methods.

#### **Details**

In a first step, a grid is created with a range from  $y_i \pm \mathrm{factor} \cdot \mathrm{range}(y_{1...n})$  with steps cuts. For each cut, the p-value is calculated for the model when  $y_i$  is shifted to that value (newobs = TRUE) or a second observation  $y_{2i}$  is added to the fixed  $y_i$  (newobs = TRUE). When the original model  $y = \beta_0 + \beta_1 x + \varepsilon$  is significant ( $p < \mathrm{alpha}$ ), there are two boundaries that result in insignificance: one decreases the slope  $\beta_1$  and the other inflates the standard error s.e.( $\beta_1$ ) in a way that  $P_t(\frac{\beta_1}{\mathrm{s.e.}(\beta_1)}, n-2) > \alpha$ . If the original model was insignificant, also two boundaries exists that either increase  $\beta_1$  or reduce s.e.( $\beta_1$ ). Often, no boundaries are found and increasing the factor grid range may alleviate this problem.

This function is quite fast ( $\sim 300 \text{ms}/10 \text{ response values}$ ), as the slope's p-value is calculated from the corr. test function of the 'psych' package, which utilizes matrix multiplication and vectorized

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pt calculation. The vector of correlation coefficients  $r_i$  from the cor function is transformed to t-values by

$$t_i = \frac{r_i \sqrt{n-2}}{\sqrt{1-r_i^2}}$$

which is equivalent to that employed in the linear regression's slope test.

### Value

A list with the following items:

X	the predictor values.
у	the response values.
pmat	the $p$ -value matrix, with length(x) columns and steps rows.
alpha	the selected $\alpha$ -level.
ySeq	the grid sequence for which the algorithm calculates $p$ -values when $y_i$ is shifted within.
model	the original 1m model.
data	the original model.frame.
eosr	the y-values of the ends of the significance region.
diff	the $\Delta$ value between $y_i$ and the nearest border of significance reversal.
closest	the (approx.) value of the nearest border of significance reversal.
newobs	should a new observation be added?

# Author(s)

Andrej-Nikolai Spiess

```
## Significant model, no new observation.
set.seed(125)
a <- 1:20
b < -5 + 0.08 * a + rnorm(length(a), 0, 1)
LM1 <- lm(b \sim a)
res1 <- lmThresh(LM1)</pre>
threshPlot(res1)
stability(res1)
## Insignificant model, no new observation.
set.seed(125)
a <- 1:20
b < -5 + 0.08 * a + rnorm(length(a), 0, 2)
LM2 \leftarrow lm(b \sim a)
res2 <- lmThresh(LM2)</pre>
threshPlot(res2)
stability(res2)
```

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```
## Significant model, new observation.
## Some significance reversal regions
## are within the prediction interval,
## e.g. 1 to 6 and 14 to 20.
set.seed(125)
a <- 1:20
b < -5 + 0.08 * a + rnorm(length(a), 0, 1)
LM3 <- lm(b \sim a)
res3 <- lmThresh(LM3, newobs = TRUE)</pre>
threshPlot(res3)
stability(res3)
## More detailed example to the above:
## a (putative) new observation within the
## prediction interval may reverse significance.
set.seed(125)
a <- 1:20
b < -5 + 0.08 * a + rnorm(length(a), 0, 1)
LM1 <- lm(b \sim a)
summary(LM1) # => p-value = 0.02688
res1 <- lmThresh(LM1, newobs = TRUE)</pre>
threshPlot(res1)
st <- stability(res1, pval = TRUE)</pre>
st$stats # => upper prediction boundary = 7.48
         # and eosr = 6.49
stabPlot(st, 1)
## reverse significance if we add a new response y_1 = 7
a <- c(1, a)
b <- c(7, b)
LM2 \leftarrow lm(b \sim a)
summary(LM2) # => p-value = 0.0767
```

shinyInfl

Initializes a Shiny App with all implemented methods of this package

#### **Description**

A comprehensive Shiny App that facilitates the import of user-supplied data and subsequent detailed testing of all implemented methods in this package. Analysis results can be exported as plots and text.

### Usage

shinyInfl()

# Arguments

None.

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#### Value

The analysis including plots and result tables.

#### Author(s)

Andrej-Nikolai Spiess

# **Examples**

```
## shinyInfl() # <= to initialize</pre>
```

simInfl

Simulates significance reversals and calculates their influence parameters

# Description

This function simulates linear regressions and stores the parameters and influence measures of all simulations that resulted in LOO significance reversal, developed for research purposes.

## Usage

```
simInfl(x = 1:10, slope = 0.02, intercept = 1, error = 0.05, nrev = 1000, ...)
```

## **Arguments**

x the x values to be supplied to 1mExact. slope the slope  $\beta_1$  to be supplied to 1mExact. intercept the intercept  $\beta_0$  to be supplied to 1mExact. error the  $\varepsilon$  value to be supplied to 1mExact. nrev the number of desired significance reversals. . . . other parameters to 1mExact and 1mInf1.

#### **Details**

Loops over an undefined number of EXACT regressions (lmExact) with incrementing random seeds, stores all models and in case of significance reversal, parameters and influence measures (lmInfl). The simulation terminates when nrev reversals are counted.

#### Value

A list with the following two items:

models the linear models of all reversals.

mat the stored matrix with the resulting parameters and influence measures for all

nrev reversals.

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#### Author(s)

Andrej-Nikolai Spiess

## **Examples**

```
## Example with slight slope, intercept = 0.5 and 10 reversals.
res -\sin Infl(x = 1:10, intercept = 0.5, slope = 0.02, error = 0.05, nrev = 10)
## Plot Cook's D versus delta-P values
## and insert common cut-off.
plot(res$mat[, "cook.d"], res$mat[, "dP"], pch = 16, cex = 0.5,
     xlab = "Cook's D", ylab = "delta-P")
thresh <- qf(0.5, 2, 8) # threshold value for Qf(0.5, npar, df)
abline(v = thresh, col = "darkred", lwd = 2)
## Plot dfbeta slope versus delta-P values
## and insert common cut-off.
plot(res$mat[, "dfb.Slope"], res$mat[, "dP"], pch = 16, cex = 0.5,
     xlab = "dfbeta Slope", ylab = "delta-P")
thresh \leftarrow 2/sqrt(10) # 2/sqrt(N)
abline(v = thresh, col = "darkred", lwd = 2)
## Plot dffits versus delta-P values
## and insert common cut-off.
plot(abs(res$mat[, "dffit"]), res$mat[, "dP"], pch = 16, cex = 0.5,
     xlab = "dffits", ylab = "delta-P")
thresh \leftarrow 2 * sqrt(2/10) # 2 * sqrt(nPar/N)
abline(v = thresh, col = "darkred", lwd = 2)
## More illustrative with more reverser samples!
## Example with slight slope, intercept = 0.5 and 10 reversals.
res -\sin Infl(x = 1:10, intercept = 0.5, slope = 0.02, error = 0.05, nrev = 200)
plot(res$mat[, "cook.d"], res$mat[, "dP"], pch = 16, cex = 0.5,
     xlab = "Cook's D", ylab = "delta-P")
thresh \leftarrow qf(0.5, 2, 8) # threshold value for Qf(0.5, npar, df)
abline(v = thresh, col = "darkred", lwd = 2)
```

stability

Calculates stability values for results of 'lmInfl', 'lmMult' and 'lmThresh'

# Description

This function calculates stability values for LOO (lmInfl), LMO (lmMult) and response value shifting/addition (lmThresh).

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#### **Usage**

stability(x, pval = FALSE, ...)

### **Arguments**

x a result of either lmInfl, lmMult or lmThresh.

pval logical. If TRUE, for lmThresh, objects an exact p-value is calculated for a future

response to reverse significance.

... other parameters, not yet implemented.

#### **Details**

For results of lmInfl:

A [0, 1]-bounded stability measure  $S = 1 - \frac{n}{N}$ , with n = number of influencers (significance reversers) and N = total number of response values.

For results of lmMult:

For each 1...max, the percentage of all resamples that did \*NOT\* result in significance reversal.

For results of lmThresh:

A [0, 1]-bounded stability measure  $S=1-\frac{n}{N}$ , with n= number of response values where one of the ends of the significance region is within the prediction interval and N= total number of response values.

If pval = TRUE, the exact p-value is calculated in the following manner:

1) Mean square error (MSE) and prediction standard error (se) are calculated from the linear model:

$$MSE = \sum_{i=1}^{n} \frac{(y_i - \hat{y}_i)^2}{n - 2} \qquad se_i = \sqrt{MSE \cdot \left(1 + \frac{1}{n} + \frac{(x_i - \bar{x}_i)^2}{\sum_{i=1}^{n} (x_i - \bar{x}_i)^2}\right)}$$

2) Upper and lower prediction intervals boundaries are calculated for each  $\hat{y}_i$ :

$$\hat{y}_i \pm Q_t(\alpha/2, n-2) \cdot \mathrm{se_i}$$

The prediction interval around  $\hat{y}_i$  is a scaled/shifted t-distribution with density function

$$P_{tss}(y, n-2) = \frac{1}{\text{se}_i} \cdot P_t \left( \frac{y - \hat{y}_i}{\text{se}_i}, n-2 \right)$$

, where  $P_t$  is the density function of the central, unit-variance t-distribution.

3) The probability of either shifting the response value (if lmThresh(..., newobs = FALSE)) or including a future response value  $y_{2i}$  (if lmThresh(..., newobs = TRUE)) to reverse the significance of the linear model is calculated as the integral between the end of the significance region (eosr) and the upper/lower  $\alpha/2$ ,  $1 - \alpha/2$  prediction interval:

$$P(\text{reverse}) = \int_{\text{eosr}}^{1-\alpha/2} P_{tss}(y, n-2) dy \quad \text{or} \quad \int_{\alpha/2}^{\text{eosr}} P_{tss}(y, n-2) dy$$

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## Value

The stability value.

#### Author(s)

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```
## See examples in 'lmInfl' and 'lmThresh'.
## The implemented strategy of calculating the
## probability of significance reversal, as explained above
## and compared to 'stabPlot'.
set.seed(125)
a <- 1:20
b < -5 + 0.08 * a + rnorm(length(a), 0, 1)
LM1 <- lm(b \sim a)
res1 <- lmThresh(LM1, newobs = TRUE)</pre>
st1 <- stability(res1, pval = TRUE)</pre>
## Let's check that the prediction interval encompasses 95%:
dt.scaled \leftarrow function(x, df, mu, s) 1/s * dt((x - mu)/s, df)
integrate(dt.scaled, lower = st1$stats[1, "lower"], st1$stats[1, "upper"],
          df = 18, mu = st1$stats[1, "fitted"], s = st1$stats[1, "se"])
## => 0.95 with absolute error < 8.4e-09
## This is the interval between "end of significance region" and upper
## prediction boundary:
integrate(dt.scaled, lower = st1$stats[1, "eosr.2"], st1$stats[1, "upper"],
          df = 18, mu = st1$stats[1, "fitted"], s = st1$stats[1, "se"])
## => 0.09264124 with absolute error < 1e-15
## We can recheck this value by P(B) - P(A):
pt.scaled <- function(x, df, mu, s) pt((x - mu)/s, df)
pA \leftarrow pt.scaled(x = st1\$stats[1, "eosr.2"], df = 18, mu = st1\$stats[1, "fitted"],
                s = st1$stats[1, "se"])
0.975 - pA
## => 0.09264124 as above
```

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